PRICE INDEX WITH GENERALIZED SPACE-TIME AUTOREGRESSIVE SEEMINGLY UNRELATED REGRESSION (GSTAR-SUR): COMPROMISE REGION AND TIME

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FORECASTING THE CONSUMER PRICE INDEX WITH GENERALIZED SPACE-TIME AUTOREGRESSIVE SEEMINGLY UNRELATED REGRESSION (GSTAR-SUR): COMPROMISE REGION AND TIME

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ABSTRACT

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Economic success will provide benefits for improving people's welfare. An important indicator to determine economic success can be seen through inflation by calculating the Consumer Price Index (CPI). CPI is a time series data that is influenced by elements between locations. The GeneralizedSpace-Time Autoregressive (GSTAR) method is a suitable method to be applied to CPI data because it involves elements of time and location [5] atiotemporal). The problem is that the GSTAR model cannot detect any correlated residuals. The GSTAR model was developed into the GSTAR-SUA model to estimate parameters with correlated residuals so produce more efficient estimates. The purpose of this study was to determine the best GSTAR-SUR model to predict the CPI of six cities in Central Java, namely Cilacap, Purwokerto, Kudus, Surakarta, Semarang 37 d Tegal. The data used is secondary data sourced from BPS Central Java Province. Based on the results of the analysis, the best model formed is the GSTAR-SUR (11)-I(1) model with an RMSE value of 6.213. Foreca 7 ng results show that the CPI value for the next 6 months will increase every month for each city.



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1. INTRODUCTION

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The Consumer Price Index (CPI) is one of the most important macroeconomic indicators [1], [2]. This is because CPI is a tool to measure inflation in an area where inflation is a controlling macroeconomic indicator that has broadly impacted serious economic indicators [3]. According to the Central Statistics Agency (BPS), CPI is an index that calculates the average price change from a collection of prices for goods and services consumed by residents or households over a certain period closely related to inflation. One of the most important Inflaton forecastings is one of the most important inputs for the decision-making process in the monetary sector. The consumer price index is also used as a basis for determining the amount of GRDP, budget planning, and other fiscal policies. According to the results of the Cost of Living Survey conducted by the Central Statistics Agency, the CPI in Indonesia generally always increases every year. In 2018 the national CPI increased by 4.14 from the previous year. Then in 2019, the national CPI increased again by 4.04 from 2018 [4].

In Indonesia, Central Java Province has a pretty good CPI. This statement can be seen from the stability of the inflation value [5], [6]. Even for the last 4 years, inflation in santral Java Province has always decreased, namely in 2017 to 2020 it decreased by 2.1 per cent. Six cities in Central Java Province are targets of the Cost of Living 19 vey, namely, Cilacap City, Purwokerto City, Kudus City, Surakarta City, Semarang City, and Tegal City. The selection of the city sample is based on the level of development in the economic field, which is relatively rapid compared to other cities [7]. These six cities will be used as the basis for this study. These six cities will be used as the basis for this study. This is because the six cities look the most prominent in economic development compared to other cities.

CPI data is included in the time series data, so it can be modeled using the ti ao series analysis method. Based on the number of variables studied, time series data can be divided into univariate time series and multivariate time series is a time series is a time series analysis that uses one variable. While the multivariate time series is a time series analysis that uses several variables in the study because it is suspected to be interrected [9]. The CPI between adjacent cities may have inter-location links [5], [6]. The linkage or correlation of the consumer price index is reflected in the relationship of dependence in meeting the needs of goods and services. Geographical conditions and limited infrastructure affect the availability of goods and services needed in an area, especially goods and services that cannot be produced alone, potentially affecting the level of costs and prices in other locations. Thus, the development of the consumer price index between cities and having a relationship with the consumer price index at the previous time also has a relationship between locations (spatial) [10].

A model the state of the space-time and involves interrelationships of events in previous times and involves interrelationships with locations in multivariate time series data are called the space-time model [8], [11], [12]. In 1980 Pfeifer and Deutsch introduced a model that combines time and location interdependence known as Space-Time Autoregressive (STAR) [13]. However, the STAR model has a weakness in the flexibility of parameters that assume that the locations have homogeneous ones, so if the locations have heterogeneous characteristics, the STAR model is not good for use [11], [12], [14]. Developed the General lized Space-Time AutoSpace-Time (GSTAR) model to address the weaknesses of the STAR model [15]. The GSTAR model is a general lization of the STAR model, which allows autoregressive (AR) parameter values to vary at each location so that the GSTAR model can be applied to heterogeneous locations. However, the study of parameter explanation using GSTAR is still limited to estimates with the Ordinary Least Square (OLS) method. So that the GSTAR model with correlated residuals wi24 produce inefficient estimators because it causes a greater error if used to do forecasting. Therefore, the Generalized Least Square (GLS) method is used to estimate parameters with correlated residuals commonly used in Seemingly Unrelated Regression (SUR) models [16], [17].

The SUR model is a pultivariate linear regression model. This model consists of several regression equations whose errors do not correlate bety 1 nobservations in an equation, but errors correlate between equations are equations as conducted by Habibie in the application of GSTAR-SUR to the number of domestic aircraft passengers at Indonesian airports. It was found that the GSTAR-SUR method gave better results than GSTAR-OLS [18]. In addition, there is also a study conducted by Septyaningrum on forecasting the number of tourists at three tourist sites in Pacitan Regency using the GSTAR-SUR method obtained from the comparison of forecasting accuracy, which shows GSTAR-SUR has a smaller RMSE value in all tourist attractions [19].

Based on the existing explanation, until now, there has been no research that specifically analyzes and forecasts the CPI of six cities that are SBH samples in Central Java using the GSTAR-SUR method.

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Therefore, this article presents the results of modeling and forecasting the consumer price index of six SBH cities in Central Java using the GSTAR-SUR time series method.

2. RESEARCH METHODS



2.1 Time Series Analysis

Time series analysis is a quantitative method to determine patterns of data in the past that are collected regularly. Time series data analysis has need purposes, including predicting future values [20]. In terms of data usage, time series analysis is divided into two, namely univariate time series and multivariate time series [8]. Univariate time series data is the result of observing one variable that has autocorrelation. In univariate time series analysis, data stability is one of ecomponents that must be considered. Analysis of univariate time series is usually modeled in several methods such as Autoregressive (AR), Moving Average (MA), Autoregressive Moving Average (ARMA), or Autoregressive Integrated Moving Average (ARIMA) [11].

A multivariate time series is a series of that consisting of several variables obtained over time and recorded sequentially according to the time of the event using a fixed time interval [21]. Multivariate time series analysis is usuall series for analysis that has more than one-time series data, so there will be many variables in the model. Similar to univariate time series, the process in multivariate time series also pays attention to data stationariness [22].

2.2 Generalized Space-Time Autoregressive Model (GSTAR)

The GSTAR model is an extens 20 of the STAR model, with its main comparison lying in the autoregressive 57 ameter. Where in the STAR model, the autoregressive 20 arameters are assumed to be the same, while in the GSTAR model, it is assumed to be heterogeneous. In matrix notation, GSTAR models with autoregressive degrees p and spatial degrees $\lambda_1, \lambda_2, ..., \lambda_p$ formulated as follows [22], [23]:

$$Z_t = \sum_{k=1}^{p} [\phi_{k0} + \phi_{k1} W] Z(t-k) + e(t)$$
 (1)



2.3 Seemingly Autoregressive Regression Model (SUR)

Seemingly Autoregressive Regression Model (SUR) is a multivariate linear regression model first troduced by Zellner [1962]. This model consists of several regression equations whose misguided ones do not correlate between observations in an equation but whose misguidedness correlates between equations. The test used to determine whether the misguided covariance variance structure is a SUR structure is a Lagrange Multiplier with a hypothesis [17]:

 $H_0 = Cov(\varepsilon_{i1}, \varepsilon_{j1}) = 0$ for all $i \neq j$ (The structure of variance and perverse covariance is heteroskedastic and there is no misguided correlation between equations).

 $H_1 = Cov(\varepsilon_{i1}, \varepsilon_{j1}) \neq 0$ for all $i \neq j$ (The structure of variance and misguided covariance is heteroskedastic and there is a misguided correlation between equations).

2.4 GSTAR-SUR Best Model Determination

The best modelling is done by looking at the smallest RMSE value of each model. Here is the RMSE formula [1], [9], [20], [24]:

$$RMSE = \sqrt{\frac{1}{M} \sum_{t=1}^{M} (Z_t - \hat{Z}_t)^2}$$
 (2)



2.5 Research Data

The data used in this study is secondary data sourced from BPS Central Java Province. The data taken is quantitative in the form of a monthly Consumer Price Index (CPI) of 6 cities in Central Java Province, namely Cilacap, Purwokerto, Kudus, Surakarta, Semarang, and Tegal, from January 2012 to December 2021, which amounts to 120 data.

No Variable Information

1 Z_{1,1} Cilacap City Monthly CPI
2 Z_{2,1} Purwokerto City Monthly CPI
3 Z_{3,1} Kudus City Monthly CPI
4 Z_{4,1} Surakarta City Monthly CPI
5 Z_{5,1} Semarang City Monthly CPI
6 Z_{6,1} Tegal City Monthly CPI

Table 1. Research Variables

2.6 Research Steps

The data analysis steps used in this study are as follows:

- 1. Determine the time series data and its location to be used in the study
- 2. Identify descriptive statistics of the CPI data of the six cities.
- 3. Examination of data stability.
- 4. Time order identification using MPACF and minimum AIC.
- 5. Determine the spatial weight used.
- 6. Perform calculation of location weight value (W1) using uniform weight and distance inverse.
- 7. Estimating p-order parameters with the GSTAR-OLS model.
- 8. Conducting a residual correlation test between locations with a covariance of the covariance of the
- Assessing the parameters of the GSTAR-SUR model by the GLS method for each location weight.
- 10. Testing the significance of GSTAR-SUR model parameters for each location weight
- 11. Testing the C17 AR-SUR model residual with residual assumption test.
- 12. Determining the best GSTAR-SUR model based on RMSE value.
- 13. Conducting GSTAR-SUR forecasting for CPI data in 6 cities in Central Java.

3. RESULTS AND DISCUSSION

3.1. Descriptive Analysis

A descriptive analysis was created to provide an overview of the data used in the study. The data used in this study is consumer price index (CPI) data for 6 cities control Java from January 2012 to December 2021, which amounts to 120 data. The picture given is the mean, standard deviation, minimum value, and maximum value according to Table 2.

Table 2. Descriptive Analysis

District/City	Mean	Standard Deviation	Minimum	Maximum
Cilacap	125.301	13.181	102.47	148.14
Purwokerto	122.529	11.882	103.33	146.97
Kudus	126.990	13.308	103.37	145.17
Surakarta	119.886	9.603	103.33	135.24
Semarang	123.317	11.708	104.41	145.29
Tegal	121.883	11.491	104.17	142.35

3.2. Correlation of CPI Data in Six Cities

The value of the correlation between regions sh49s how much the relationship between regions is and another. The correlation values between the six cities can be seen in Table 3.

Table 3. Correlation of CPI Data in Six Cities

	35					
City	Cilacap	Purwokerto	Kudus	Surakarta	Semarang	Tegal
Cilacap	1	0.9804	0.9818	0.9917	0.9883	0.9702
Purwokerto	0.9804	1	0.9356	0.9646	0.9951	0.9897
Kudus	0.9818	0.9356	1	0.9896	0.9519	0.9268
Surakarta	0.9917	0.9646	0.9896	1	0.9779	0.9607
Semarang	0.9883	0.9951	0.9519	0.9779	1	0.9936
Tegal	0.9702	0.9897	0.9268	0.9607	0.9936	1

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Based on Table 3, it can be seen that the CPI between six cities in Central Java has a high correlation value. This indicates that there is a linkage of CPI data in the same time order and shows that the CPI between adjacent locations has a high relationship with each other.

3.3. Stationary Test



In time series data modeling, it must meet two assumptions, namely the data must be stationary to the mean and variance. Stationariness tests against variance can be seen through box-cox plot results, while stationary tests against mean can be seen from Augmented Dickey-Fuller (ADF) tests.

3.3.1 Stationary Test of the Mean



Time series data sensitivity testing against the mean can be performed with the Augmented Dickey-Fuller (ADF) test. The hypotheses used in the ADF test are as follows:

 H_0 : $\delta = 1$ (non-stationary data)

 $\overline{H_1}$: $\delta < 1$ (stationary data)

 H_0 is accepted if the *p-value* > $\alpha = 0.05$.

Table 4. ADF Test

City	p-value	α	10 formation
Cilacap	0.4439		Non-stationary
Purwokerto	0.3917		Non-stationary
Kudus	0.544	0.05	Non-stationary
Surakarta	0.4313	0.05	Non-stationary
Semarang	0.3816		Non-stationary
Tegal	0.4101		Non-stationary

Based on Table $\frac{53}{4}$, it can be seen that the CPI data of six SBH cities in Central Java is not yet stationary. This is evidenced by the p-value of each city 16 hich is > 0.05, which means that H₀ is accepted. Therefore it is necessary to be different from stationing the data. The results of the ADF test after differencing are found in Table 5 as follows:

Table 5. ADF Test after differencing

			0
City	p value	α	Information
Cilacap	0.01		stationary
Purwokerto	0.01		stationary
Kudus	0.01	0.05	stat ionary
Surakarta	0.01	0.05	stationary
Semarang	0.01		stationary
Tegal	0.01		stationary

After differencing, it can be seen in Table 5 that the CPI data of six cities have been stationary. This is evidenced by the p-value of each city < 0.05, meaning that H_0 is rejected so that it can be said that the data is stationary.

3.3.2. Stationary Test of the Variance

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Time series data sensitivity testing against variance can be done with the Box-Cox test. Data is said to be stationary to a variance if its rounded value is worth 1 (λ =1). Here are the Box-Cox test results for six cities:

Table	6. Box-Cox	Test
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City	Rounded value	10 formation
Cilacap	-2.00	Non-stationary
Purwokerto	-3.00	Non-stationary
Kudus	-1.00	Non-stationary
Surakarta	-2.00	Non-stationary
Semarang	-1.00	Non-stationary
Tegal	-0.50	Non-stationary

so sed on the Box-Cox test results for the six locations, the results were obtained that the six locations were not stationary to the variance because the rounded value was $\neq 1$ and had different lambda estimate values for each location. If a transformation is performed, then the transformation used is different according to the lambda estimate for each location. Therefore, there is no need for transformations and the data is considered stationary against variance.

3.4. Test the Residual Correlation between Locations

Before the GS32 $\ R$ (1₁)-I(1) model was modeled using Seemingly Unrelated Regression (SUR), a Lagrange Multiplier test was first carried out to determine whether there was a residual correlation between locations in the GSTAR (1₁)-I(1) model. Here's the Lagrange Multiplier test.

Hypothesis:

$$H_0 = Cov(\varepsilon_{i1}, \varepsilon_{j1}) = 0$$
 for all $i \neq j$

(The structure of variance and misguided covariance is heteroscedastic, and there is no residual correlation between equations).

$$H_1 = Cov(\varepsilon_{i1}, \varepsilon_{i1}) \neq 0$$
 for all $i \neq j$

(The structure of variance and misguided covariance is heteroscedastic, and there is a residual correlation between equations).

Significance: $\alpha = 5\%$

Test criteria: $\lambda_{LM} > \chi^2_{0,05;15} = 24.996$

Test statistics: $\lambda_{LM} = T \sum_{i=2}^{N} \sum_{j=1}^{i-1} r_{ij}^2$

Here are the statistical values of the Lagrange Multiplier test for each location weight:

1. GSTAR model (1₁)-I(1) with Uniform Location Weights

$$\begin{split} \lambda_{LM} &= 118(0.919^2 + 0.932^2 + 0.948^2 + 0.709^2 + 0.95^2 + 0.922^2 + 0.929^2 + 0.682^2 + 0.935^2 + 0.975^2 + 0.690^2 + 0.910^2 + 0.709^2 + 0.940^2 + 0.705^2) \\ \lambda_{LM} &= 1322,\!506 \end{split}$$

Since $\lambda_{LM} = 1322,506 > \chi^2_{0,05;15} = 24.996$ then it can be concluded that H_0 is rejected so that it can be stated that there is a residual correlation between locations.

2. GSTAR model (1₁)-I(1) with Distance Inverse Location Weights

$$\begin{array}{l} \lambda_{LM} = 118(0.961^2 + 0.822^2 + 0.821^2 + 0.826^2 + 0.786^2 + 0.667^2 + 0.664^2 + 0.676^2 + 0.630^2 + 0.980^2 + 0.945^2 + 0.907^2 + 0.967^2 + 0.934^2 + 0.975^2) \\ \lambda_{LM} = 1268.161 \end{array}$$

Since $\lambda_{LM} = 1268.161 > \chi^2_{0.05;15} = 24.996$ then it can be concluded that $\overline{H_0}$ is rejected so that it can be stated that there is a residual correlation between locations.

3.5. GSTAR-SUR Model Parameter Estimation by GLS Method

After the Lagrange Multiplier test is carried out, the residual correlated model is processed using the GSTAR-SUR model. Parameter estimation in the GSTAR-SUR model can be done by the Generalized Least Square (GLS) method. This method's assessment of significant parameter estimates is the same as in tags GSTAR model. That is, the value of parameter estimates is said to be significant if $|t_{value}| > t_{table}$. However, in this study, insignificant parameters were not eliminated with the purpose of weighting from each location remaining. Here is an estimate of the parameters of the GSTAR-SUR model of the six locations based on each weighting:

3.5.1 Estimation of GSTAR-SUR (11)-I(1) Model Parameters with Uniform Weights

The result of the calculation of the estimated parameters of the GSTAR-SUR (11)-I(1) model using uniform weights can be seen in Table 7 as follows:

Table 7. Estimation of GSTAR-SUR (11)-I(1) Model Parameters with Uniform Weights

Parameter	Estimation Values	t-value	t-table	Information
ϕ_{10}^1	-0.1922	-1.264		Insignificant
ϕ_{10}^{2}	-0.4502	-5.232		Significant
ϕ_{10}^{3}	0.0392	0.344		Insignificant
ϕ_{10}^4	0.2069	0.922		Insignificant
ϕ_{10}^{5}	-1.3794	-6.148		Significant
ϕ_{10}^6	-0.2817	-1.569	1.98	Insignificant
ϕ_{11}^1	0.1755	1.003	1.96	Insignificant
ϕ_{11}^{2}	0.4106	4.572		Significant
ϕ_{11}^{3}	-0.0736	-0.596		Insignificant
ϕ_{11}^4	-0.1956	-1.068		Insignificant
ϕ_{11}^{5}	1.3665	5.987		Significant
ϕ_{11}^6	0.2381	1.357		Insignificant

3.5.2 Estimation of GSTAR-SUR (11)-I(1) Model Parameters with Distance Inverse Weights

The results of the 9 Iculation of the estimated parameters of the GSTAR-SUR (1₁)-I(1) model using distance inverse weights can be seen in Table 8 as follows:

Table 8. Estimation of GSTAR-SUR (11)-I(1) Model Parameters with Distance Inverse Weights

Parameter	Estimation Values	t-value	t-table	Information 33
ϕ_{10}^1	0.1264	1.045		Insignificant
ϕ_{10}^{2}	-0.4537	-5.146		Significant
ϕ_{10}^{3}	0.0028	0.024		Insignificant
ϕ_{10}^4	0.1286	0.488		Insignificant
ϕ_{10}^{5}	-0.0387	-0.216	1.98	Insignificant
ϕ_{10}^6	-0.0269	-0.155		Insignificant
ϕ_{11}^{1}	-0.1843	-1.348		Insignificant
ϕ_{11}^{2}	0.4008	4.525		Significant
ϕ_{11}^{3}	-0.0337	-0.265		Insignificant
ϕ_{11}^4	-0.1396	-0.648		Insignificant
ϕ_{11}^5	0.0091	0.050		Insignificant
ϕ_{11}^6	-0.0085	-0.051		Insignificant

3.6. GSTAR-SUR Best Model Determination (1₁)-I(1)

After obtaining the GSTAR-SUR model and testing the feasibility of the model, the calculation of the accuracy of the model is carried out to obtain the best model. Determining the best mode 56 done by looking at the RMSE value of each model. Later 22 model with the smallest RMSE is declared the best model. The RMSE calculation results of each model can be seen in Table 9 as follows:

Table 9. RMSE Model Value

7723322323							
Weights			(City			- Mean of RMSE
weights	Cilacap	Purwokerto	Kudus	Surakarta	Semarang	Tegal	Weall of KMSE
Seragam	7.279	6.319	6.530	5.108	6.552	6.029	6.303
Invers Jarak	7.054	6.374	6.542	5.106	6.199	6.001	6.213

The results of the calcul 26 on of the RMSE value can be seen that the average RMSE value of the GSTAR-SUR model (1₁)-I value using the inverse distance weight has a smaller value compared to other weights, which is 6.213. So it can be concluded that the GSTAR-SUR (1₁)-I(1) model with inverse distance weighting is the best model.

3.7. Forecasting CPI Data for Each City Using the Best Model

The best model that has been obtained is the GSTAR SUR (11)-I(1) distance inverse weight. This model is then used as a forecast to get predictions of Consumer Price Index (CPI) data or six cities in Central Java, namely, Cilacap, Purwokerto, Kudus, Surakarta, Semara 22 and Tegal. Data forecasting is carried out from January 2022 to December 2022. The forecasting results can be seen in Table 10 as follows:

Table 10. CPI Forecasting Results

Month	City					
Month	Cilacap	Purwokerto	Kudus	Surakarta	Semarang	Tegal
January-22	107.072	107.918	106.839	108.082	108.121	108.591
February-22	107.936	108.696	107.359	108.857	108.753	109.292
March-22	108.800	109.470	107.878	109.633	109.385	109.994
April-22	109.664	110.245	108.398	110.408	110.017	110.695
Mei-22	110.528	111.020	108.917	111.183	110.648	111.396
June-22	111.393	111.795	109.437	111.959	111.280	112.098
July-22	112.257	112.570	109.956	112.734	111.912	112.799
Agustus-22	113.121	113.345	110.475	113.510	112.544	113.500
September-22	113.985	114.120	110.995	114.285	113.175	114.202
Oktober-22	114.849	114.895	111.514	115.060	113.807	114.903
November-22	115.713	115.670	112.034	115.836	114.439	115.604
December-22	116.577	116.446	112.553	116.611	115.071	116.305

Based on Table 10, it is found that the forecasting value of the CPI of the six cities for the next 12 months will increase every month. However, the increase in CPI that occurred was still quite stable. This can be seen from the absence of a significant spike from month one to the next.

4. CONCLUSIONS

Based on the results of the analysis and discussion that has been carried out, it can be concluded that based on the RMSE value that has been obtained, the GSTAR-SUR (1₁)-I(1) model with distance inverse weighting is the best model because it has an RMSE value of 6.213, where this value is the smallest RMSE value compared to other weights. The results of CPI forecasting from the GSTAR-SUR (1₁)-I(1) model with distance inverse weighting for the next 12 months, namely January 2022 to December 2022, found that the CPI value for the six cities will increase every month.

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